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A Hybrid Method to Improve Forecasting Accuracy - An Introduction of a Day of the Week Index

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Abstract— Sales forecasting is inevitable for a shop owner to manage the shop smoothly. In this paper, we propose a new method to improve forecasting accuracy and confirm them by the numerical example. Focusing that the equation of exponential smoothing method(ESM) is equivalent to (1,1) order ARMA model equation, a new method of estimation of smoothing constant in exponential smoothing method is proposed before by us which satisfies minimum variance of forecasting error. Generally, smoothing constant is selected arbitrarily. But in this paper, we utilize above stated theoretical solution. Firstly, we make estimation of ARMA model parameter and then estimate smoothing constants. Thus theoretical solution is derived in a simple way and it may be utilized in various fields. Combining the trend removing method with this method, we aim to improve forecasting accuracy. Furthermore, “a day of the week index” is newly introduced for the daily data and we have obtained good result.

The effectiveness of this method should be examined in various cases.

Index Terms—Exponential Smoothing Method, Forecasting, Minimum Variance, Trend.

I. INTRODUCTION

Many methods for time series analysis have been presented such as Autoregressive model (AR Model), Autoregressive Moving Average Model (ARMA Model) and Exponential Smoothing Method (ESM) (Box Jenkins [1]), (R.G.Brown[11]), (Tokumaru et al.[3]), (Kobayashi[7]). Among these, ESM is said to be a practical simple method.

For this method, various improving method such as adding compensating item for time lag, coping with the time series with trend (Peter [10]), utilizing Kalman Filter (Maeda [4]), Bayes Forecasting (M.West et al. [8]), adaptive ESM (Steinar [13]), exponentially weighted Moving Averages with irregular updating periods (F.R.Johnston [2]), making averages of forecasts using plural method (Spyros [12]) are presented. For example, Maeda[4] calculated smoothing constant in relationship with S/N ratio under the assumption that the observation noise was added to the system. But he had to calculate under supposed noise because he could not grasp observation noise. It can be said that it doesn't pursue optimum solution from the very data themselves which should be derived by those estimation. Ishii [9] pointed out that the optimal smoothing constant was the solution of infinite order equation, but he didn't show analytical solution. Based on these facts, we proposed a new method of estimation of smoothing constant in ESM before (Takeyasu et al. [6]). Focusing that the equation of ESM is equivalent to (1,1) order ARMA model equation, a new method of estimation of smoothing constant in ESM was derived.

In this paper, utilizing above stated method, a revised forecasting method is proposed. In making forecast such as sales data, trend removing method is devised. Trend removing by the combination of linear and 2nd order non-linear function and 3rd order non-linear function is executed to the original sales data of a shop. The weights for these functions are varied by 0.01 increment and optimal weights are searched.

“a day of the week index (DWI)” is newly introduced for the daily sales data and a day of the week trend is removed. Theoretical solution of smoothing constant of ESM is calculated for both of the DWI trend removing data and the non DWI trend removing data. Then forecasting is executed on these data. This is a revised forecasting method. Variance of forecasting error of this newly proposed method is assumed to be less than those of previously proposed method. The rest of the paper is organized as follows. In section 2, ESM is stated by ARMA model and estimation method of smoothing constant is derived using ARMA model identification. The combination of linear



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and non-linear function is introduced for trend removing in section 3. “a day of the week index (DWI)” is newly introduced in section 4. Forecasting is executed in section 5, and estimation accuracy is examined.

II. DESCRIPTION OF ESM USING ARMA MODEL

In ESM, forecasting at time $t+1$ is stated in the following equation.

$$\begin{aligned}\hat{x}_{t+1} &= \hat{x}_t + \alpha(x_t - \hat{x}_t) \\ &= \alpha x_t + (1 - \alpha)\hat{x}_t\end{aligned}\tag{1}$$

Here,

\hat{x}_{t+1} : Forecasting at $t+1$

x_t : Realized value at t

α : Smoothing constant ($0 < \alpha < 1$)

(1) is re-stated as:

$$\hat{x}_{t+1} = \sum_{i=0}^{\infty} \alpha(1-\alpha)^i x_{t-i}\tag{2}$$

By the way, we consider the following (1,1) order ARMA model.

$$x_t - x_{t-1} = e_t - \beta e_{t-1}\tag{3}$$

Generally, (p, q) order ARMA model is stated as:

$$x_t + \sum_{i=1}^p a_i x_{t-i} = e_t + \sum_{j=1}^q b_j e_{t-j}\tag{4}$$

Here,

$\{x_t\}$: Sample process of Stationary Ergodic Gaussian Process $x(t) \quad t=1,2,\dots,N,\dots$

$\{e_t\}$: Gaussian White Noise with 0 mean σ_e^2 variance

MA process in (4) is supposed to satisfy convertibility condition.

Utilizing the relation that:

$$E[e_t | e_{t-1}, e_{t-2}, \dots] = 0$$

we get the following equation from (3).

$$\hat{x}_t = x_{t-1} - \beta e_{t-1}\tag{5}$$

Operating this scheme on $t+1$, we finally get:

$$\begin{aligned}\hat{x}_{t+1} &= \hat{x}_t + (1 - \beta)e_t \\ &= \hat{x}_t + (1 - \beta)(x_t - \hat{x}_t)\end{aligned}\tag{6}$$

If we set $1 - \beta = \alpha$, the above equation is the same with (1), i.e., equation of ESM is equivalent to (1,1) order ARMA model, or is said to be (0,1,1) order ARIMA model because 1st order AR parameter is -1 (Box Jenkins [1]), (Tokumaru et al.[3]).

Comparing with (3) and (4), we obtain:

$$\left\{ \begin{array}{l} a_1 = -1 \\ b_1 = -\beta \end{array} \right.$$

From (1), (6),



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$$\alpha = 1 - \beta$$

Therefore, we get:

$$\left\{ \begin{array}{l} a_1 = -1 \\ b_1 = -\beta = \alpha - 1 \end{array} \right. \quad (7)$$

From above, we can get estimation of smoothing constant after we identify the parameter of MA part of ARMA model. But, generally MA part of ARMA model become non-linear equations which are described below.

Let (4) be:

$$\left\{ \begin{array}{l} \tilde{x}_t = x_t + \sum_{i=1}^p a_i x_{t-i} \\ \tilde{x}_t = e_t + \sum_{j=1}^q b_j e_{t-j} \end{array} \right. \quad (8)$$

$$\left\{ \begin{array}{l} \tilde{x}_t = e_t + \sum_{j=1}^q b_j e_{t-j} \end{array} \right. \quad (9)$$

We express the autocorrelation function of \tilde{x}_t as \tilde{r}_k and from (8), (9), we get the following non-linear equations which are well known^[3].

$$\left. \begin{array}{l} \tilde{r}_k = \sigma_e^2 \sum_{j=0}^{q-k} b_j b_{k+j} \quad (k \leq q) \\ 0 \quad (k \geq q+1) \\ \tilde{r}_0 = \sigma_e^2 \sum_{j=0}^q b_j^2 \end{array} \right\} \quad (10)$$

For these equations, recursive algorithm has been developed. In this paper, parameter to be estimated is only b_1 , so it can be solved in the following way.

From (3) (4) (7) (10), we get:

$$\left. \begin{array}{l} q = 1 \\ a_1 = -1 \\ b_1 = -\beta = \alpha - 1 \\ \tilde{r}_0 = (1 + b_1^2) \sigma_e^2 \\ \tilde{r}_1 = b_1 \sigma_e^2 \end{array} \right\} \quad (11)$$

If we set:

$$\rho_k = \frac{\tilde{r}_k}{\tilde{r}_0} \quad (12)$$

the following equation is derived.

$$\rho_1 = \frac{b_1}{1 + b_1^2} \quad (13)$$

We can get b_1 as follows.

$$b_1 = \frac{1 \pm \sqrt{1 - 4\rho_1^2}}{2\rho_1} \quad (14)$$

In order to have real roots, ρ_1 must satisfy:

$$|\rho_1| \leq \frac{1}{2} \quad (15)$$

From invertibility condition, b_1 must satisfy:



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$$|b_1| < 1$$

From (13), using the next relation,

$$(1 - b_1)^2 \geq 0$$

$$(1 + b_1)^2 \geq 0$$

(15) always holds.

As

$$\alpha = b_1 + 1$$

b_1 is within the range of:

$$-1 < b_1 < 0$$

Finally we get:

$$\left. \begin{aligned} b_1 &= \frac{1 - \sqrt{1 - 4\rho_1^2}}{2\rho_1} \\ \alpha &= \frac{1 + 2\rho_1 - \sqrt{1 - 4\rho_1^2}}{2\rho_1} \end{aligned} \right\} \quad (16)$$

Which satisfy above condition? Thus we can obtain a theoretical solution by a simple way.

Here ρ_1 must satisfy:

$$-\frac{1}{2} < \rho_1 < 0 \quad (17)$$

in order to satisfy $0 < \alpha < 1$.

Focusing on the idea that the equation of ESM is equivalent to (1,1) order ARMA model equation, we can estimate smoothing constant after estimating ARMA model parameter.

It can be estimated only by calculating 0th and 1st order autocorrelation function.

III. TREND REMOVAL METHOD

As trend removal method, we describe the combination of linear and non-linear function.

[1] Linear function

We set:

$$y = a_1x + b_1 \quad (18)$$

as a linear function.

[2] Non-linear function

We set:

$$y = a_2x^2 + b_2x + c_2 \quad (19)$$

$$y = a_3x^3 + b_3x^2 + c_3x + d_3 \quad (20)$$

as a 2nd and a 3rd order non-linear function.

[3] The combination of linear and non-linear function

We set:

$$y = \alpha_1(a_1x + b_1) + \alpha_2(a_2x^2 + b_2x + c_2) + \alpha_3(a_3x^3 + b_3x^2 + c_3x + d_3) \quad (21)$$

$$0 \leq \alpha_1 \leq 1, 0 \leq \alpha_2 \leq 1, 0 \leq \alpha_3 \leq 1 \quad (22)$$

$$\alpha_1 + \alpha_2 + \alpha_3 = 1$$

as the combination of linear and 2nd order non-linear and 3rd order non-linear function.



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Trend is removed by dividing the data by (21).

IV. A DAY OF THE WEEK INDEX

“a day of the week index (DWI)” is newly introduced for the daily data. It is often seen that sales increase at Monday or Friday if the shop is closed on Saturday and Sunday. Therefore a day of the week index may be considered and the forecasting accuracy would improve after we identify the “a day of the week index” and utilize them. This time in this paper, the data we handle consist by Monday through Friday, we calculate DWI_j ($j = 1, \dots, 5$) for Monday through Friday. For example, if there is the daily data of L weeks as stated below:

$$\{x_{ij}\} (i = 1, \dots, L) (j = 1, \dots, 5)$$

where $x_{ij} \in R$ in which L means the number of weeks (Here $L = 8$), i means the order of weeks (i -th week), j means the order in a week (j -th order in a week; for example $j = 1$: Monday, $j = 5$: Friday) and x_{ij} is a shipping data. Then, DWI_j is calculated as follows.

$$DWI_j = \frac{\frac{1}{L} \sum_{i=1}^L x_{ij}}{\frac{1}{L} \cdot \frac{1}{5} \sum_{i=1}^L \sum_{j=1}^5 x_{ij}} \quad (23)$$

DWI trend removal is executed by dividing the data by (23). Numerical examples both of DWI removal case and non-removal case are discussed in section 5.

V. FORECASTING THE SALES DATA

A. Analysis Procedure

The sales data of a shop for 6 cases from October 1, 2009 (Thurs.) to December 9, 2009 (Wed.) are analyzed. First of all, graphical charts of these time series data are exhibited in Figure 5-1 to 5-6.



Fig. 5-1. Sales data of Food A

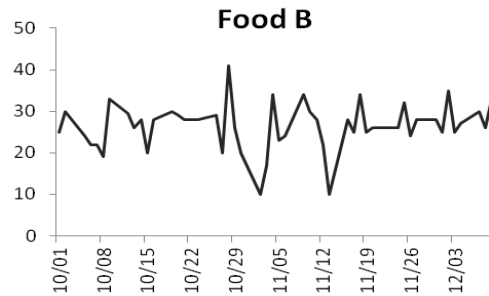


Fig. 5-2. Sales data of Food B

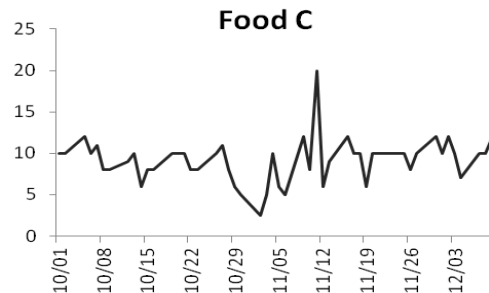


Fig. 5-3. Sales data of Food C



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Fig. 5-4. Sales data of Food D

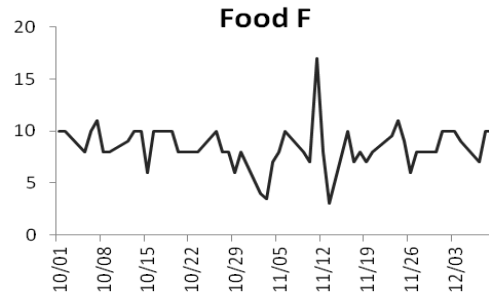


Fig. 5-6. Sales data of Food F

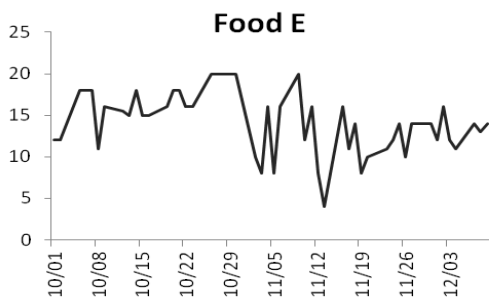


Fig. 5-5. Sales data of Food E

Analysis procedure is as follows. There are 50 daily sales data for each case. We use 40 data (1 to 40) and remove trend by the method stated in section 3. Then we calculate a day of the week index (DWI) by the method stated in section 4. After removing DWI trend, the method stated in section 2 is applied and Exponential Smoothing Constant with minimum variance of forecasting error is estimated. Then 1 step forecast is executed. Thus, data is shifted to 2nd to 41st and the forecast for 42nd data is executed consecutively, which finally reaches forecast of 50th data. To examine the accuracy of forecasting, variance of forecasting error is calculated for the data of 41st to 50th data. Final forecasting data is obtained by multiplying DWI and trend.

Forecasting error is expressed as:

$$\varepsilon_i = \hat{x}_i - x_i \quad (24)$$

$$\bar{\varepsilon} = \frac{1}{N} \sum_{i=1}^N \varepsilon_i \quad (25)$$

Variance of forecasting error is calculated by:

$$\sigma_{\varepsilon}^2 = \frac{1}{N-1} \sum_{i=1}^N (\varepsilon_i - \bar{\varepsilon})^2 \quad (26)$$

In this paper, we examine the two cases stated in Table 5-1.

Table 5-1. The combination of the case of trend removal and DWI trend removal

Case	Trend	DWI trend
Case1	Removal	Removal
Case2	Removal	Non removal

B. Trend Removing

Trend is removed by dividing original data by (21). Here, the weight of and are shifted by 0.01 increment in (21) which satisfy the equation (22). The best solution is selected which minimizes the variance of forecasting error. Estimation results of coefficient of (18), (19) and (20) are exhibited in Table 5-2. Data are fitted to (18), (19) and



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(20), and using the least square method, parameters of (18), (19) and (20) are estimated. Estimation results of weights of (21) are exhibited in Table 5-3. The weighting parameters are selected so as to minimize the variance of forecasting error.

Table 5-2. Coefficient of (18),(19) and (20)

Food	1 st		2 nd			3 rd			
	a_1	b_1	a_2	b_2	c_2	a_3	b_3	c_3	d_3
A	-0.476	52.562	-0.006	-0.251	50.982	0.0012	-0.0818	1.016	46.388
B	0.009	25.794	0.002	-0.065	26.317	0.0009	-0.0550	0.878	22.897
C	0.003	8.885	0.007	-0.282	10.879	0.0157	-0.428	11.407	0.0157
D	-0.272	27.285	0.008	-0.605	29.615	0.0002	-0.0058	-0.374	28.776
E	-0.136	17.333	-0.010	0.292	14.332	0.0008	-0.0590	1.099	11.408
F	-0.025	8.931	0.005	-0.241	10.446	0.0001	-0.0032	-0.100	9.933

Table 5-3. Weights of (21)

Food	Case	α_1	α_2	α_3
A	Case1	1.00	0.00	0.00
	Case2	0.77	0.23	0.00
B	Case1	1.00	0.00	0.00
	Case2	1.00	0.00	0.00
C	Case1	0.00	0.93	0.07
	Case2	1.00	0.00	0.00
D	Case1	1.00	0.00	0.00
	Case2	1.00	0.00	0.00
E	Case1	1.00	0.00	0.00
	Case2	0.95	0.05	0.00
F	Case1	1.00	0.00	0.00
	Case2	0.27	0.73	0.00

As a result, we can observe the following three patterns.

① Selected liner model:

Food A Case1, Food B Case1/Case2, Food C Case2, Food D Case1/Case2, Food E Case1, Food F Case 1

② Selected 1st+2nd order model:

Food A Case2, Food F Case2

③ Selected 2nd +3rd order model:

Food C Case1

Graphical charts of trend are exhibited in Figure 5-7 to 5-12.



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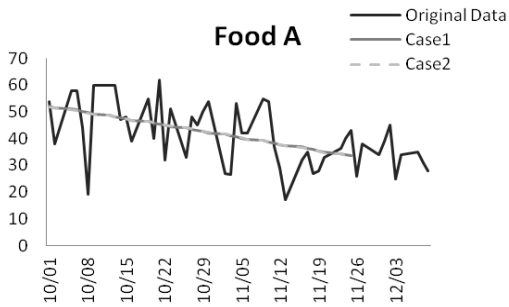


Fig. 5-7. Trend of Food A

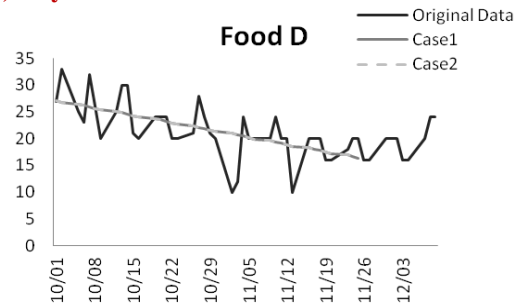


Fig. 5-10. Trend of Food D

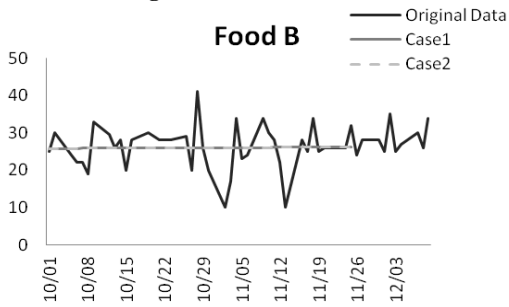


Fig. 5-8. Trend of Food B

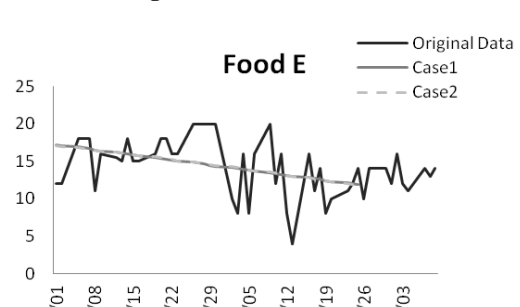


Fig. 5-11. Trend of Food E



Fig. 5-9. Trend of Food C

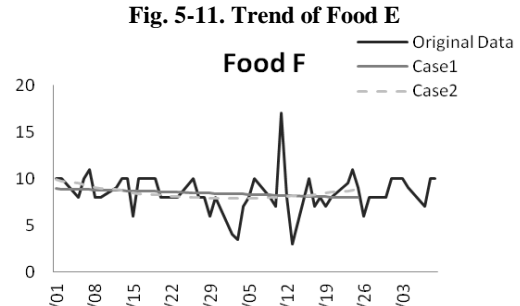


Fig. 5-12. Trend of Food F

C. Removing Trend by DWI

After removing trend, a day of the week index is calculated by the method stated in 4. Calculation result for 1st to 40th data is exhibited in Table 5-4.

Table 5-4. a day of the week index

Food	Case	a day of the week index				
		Thurs.	Fri.	Mon.	Tues.	Wed.
A	Case1	0.8655	0.9596	1.0351	1.0676	1.0722
B	Case1	0.9049	0.9583	1.0125	0.9375	1.1868
C	Case1	0.8099	0.8748	1.0813	1.0383	1.1957
D	Case1	0.9605	0.8928	0.9415	1.0618	1.1434
E	Case1	0.8237	0.9247	1.0918	0.9831	1.1767
F	Case1	0.8989	0.9584	1.0178	0.9593	1.1656

D. Estimation of Smoothing Constant with Minimum Variance of Forecasting Error

After removing DWI trend, Smoothing Constant with minimum variance of forecasting error is estimated utilizing



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(16). There are cases that we cannot obtain a theoretical solution because they do not satisfy the condition of (17). In those cases, Smoothing Constant with minimum variance of forecasting error is derived by shifting variable from 0.01 to 0.99 with 0.01 interval. Calculation result for 1st to 40th data is exhibited in Table 5-5.

Table 5-5. Estimated Smoothing Constant with Minimum Variance

Food	Case	ρ_1	α
A	Case1	-0.42800	0.43571
	Case2	-0.41590	0.46510
B	Case1	-0.27781	0.69662
	Case2	-0.38555	0.52888
C	Case1	-0.76572 (Does not satisfy (17))	0.05000
	Case2	-0.72664 (Does not satisfy (17))	0.03000
D	Case1	-0.38280	0.53411
	Case2	-0.23509	0.75025
E	Case1	-0.33001	0.62311
	Case2	-0.42630	0.44002
F	Case1	-0.30070	0.66569
	Case2	-0.00003	0.99997

E. Forecasting and Variance of Forecasting Error

Utilizing smoothing constant estimated in the previous section, forecasting is executed for the data of 41st to 50th data. Final forecasting data is obtained by multiplying DWI and trend. Variance of forecasting error is calculated by (26). Forecasting results are exhibited in Figure 5-13 to 5-18



Fig. 5-13. Forecasting Results of Food A



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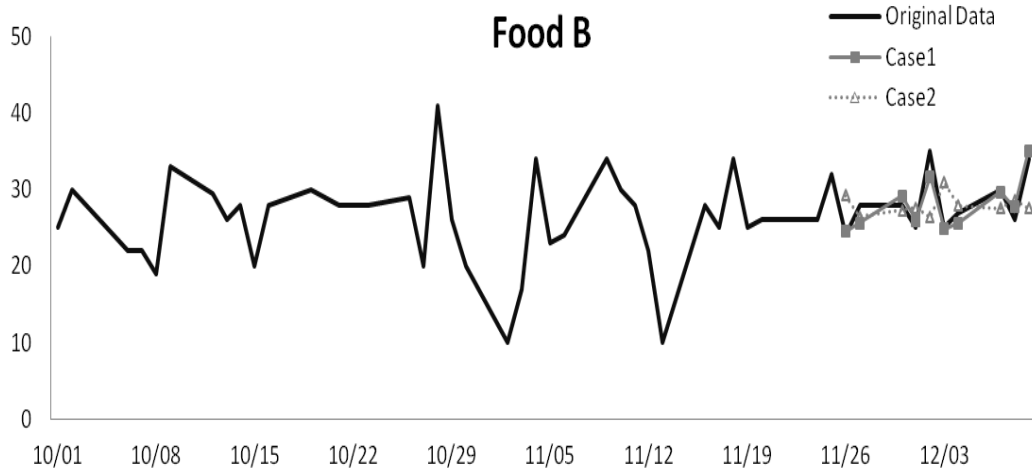


Fig. 5-14 Forecasting Results of Food B

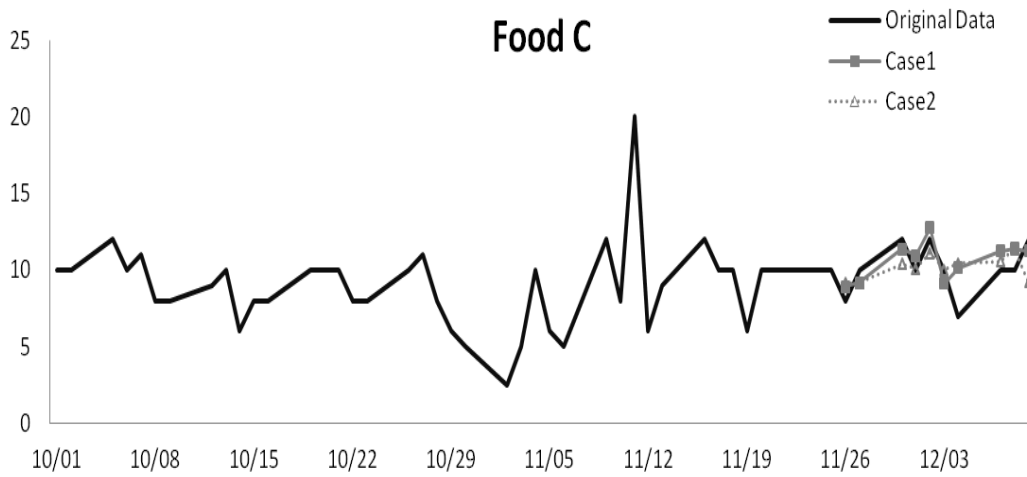


Fig. 5-15. Forecasting Results of Food C

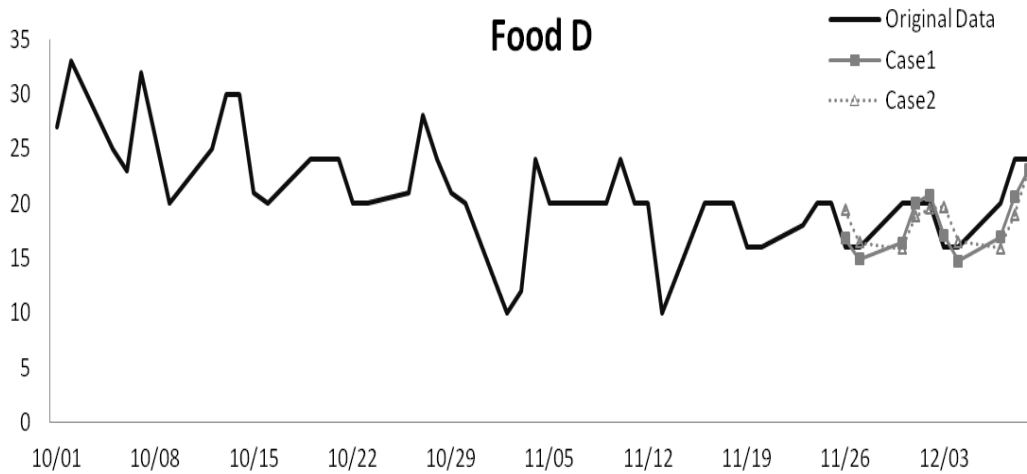


Fig. 5-16. Forecasting Results of Food D

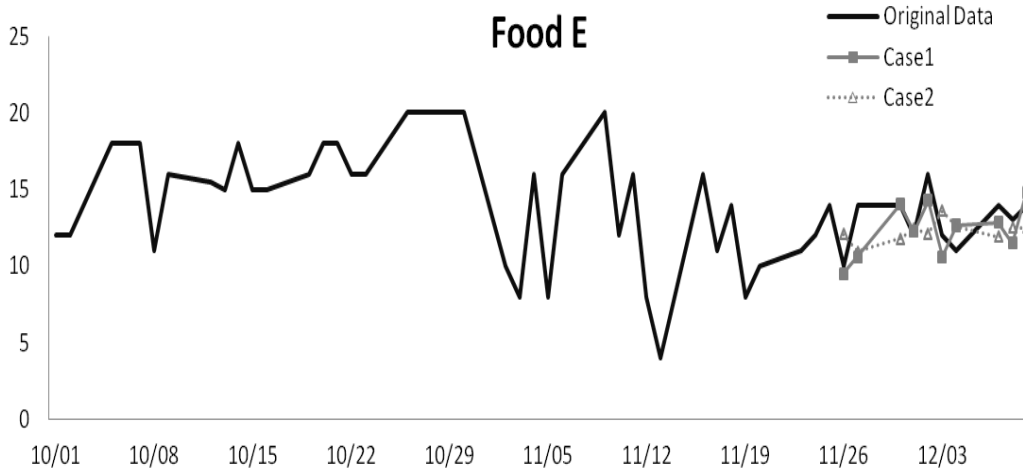


Fig. 5-17. Forecasting Results of Food E

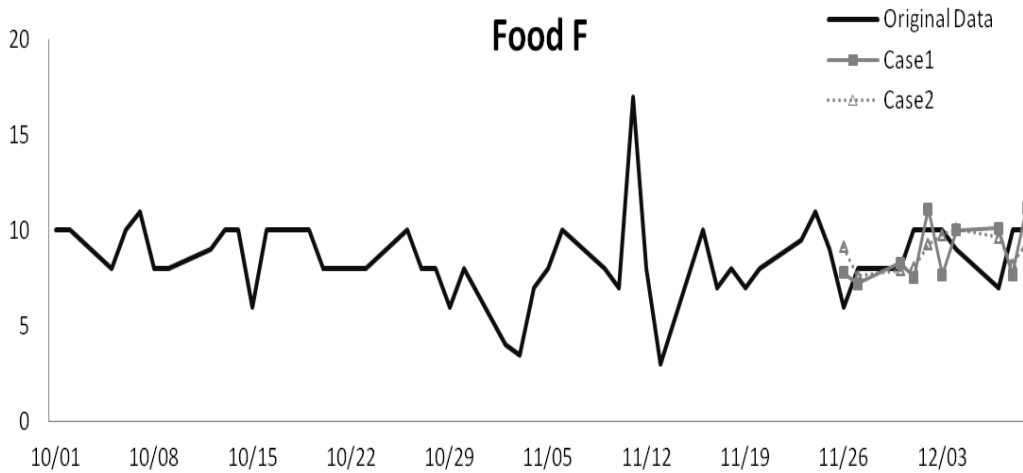


Fig. 5-18. Forecasting Results of Food F

Variance of forecasting error is exhibited in Table 5-6.

Table 5-6. Variance of Forecasting Error

Food	Case	Variance of Forecasting Error
A	Case1	29.4 *
	Case2	64.8
B	Case1	2.8 *
	Case2	22.5
C	Case1	1.7 *
	Case2	3.0
D	Case1	3.2 *
	Case2	9.0
E	Case1	2.1 *
	Case2	4.5
F	Case1	3.8
	Case2	2.9 *



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Now, we calculate the following Fluctuation Index for the comparison of forecasting accuracy by each goods.

$$FI = \frac{\sigma_{N_0/N_1}}{\bar{x}_{N_0/N_1}} \quad (27)$$

where

FI: Fluctuation Index

σ_{N_0/N_1} : Standard Deviation of forecasting error at forecasting period N_0 to N_1

\bar{x}_{N_0/N_1} :mean of realized value at forecasting period N_0 to N_1

FI is exhibited in Table 5-7 and Figure 5-19.

Table 5-7. FI

Food	Case	FI
A	Case1	16
	Case2	24
B	Case1	6
	Case2	17
C	Case1	13
	Case2	17
D	Case1	9
	Case2	16
E	Case1	11
	Case2	16
F	Case1	22
	Case2	19

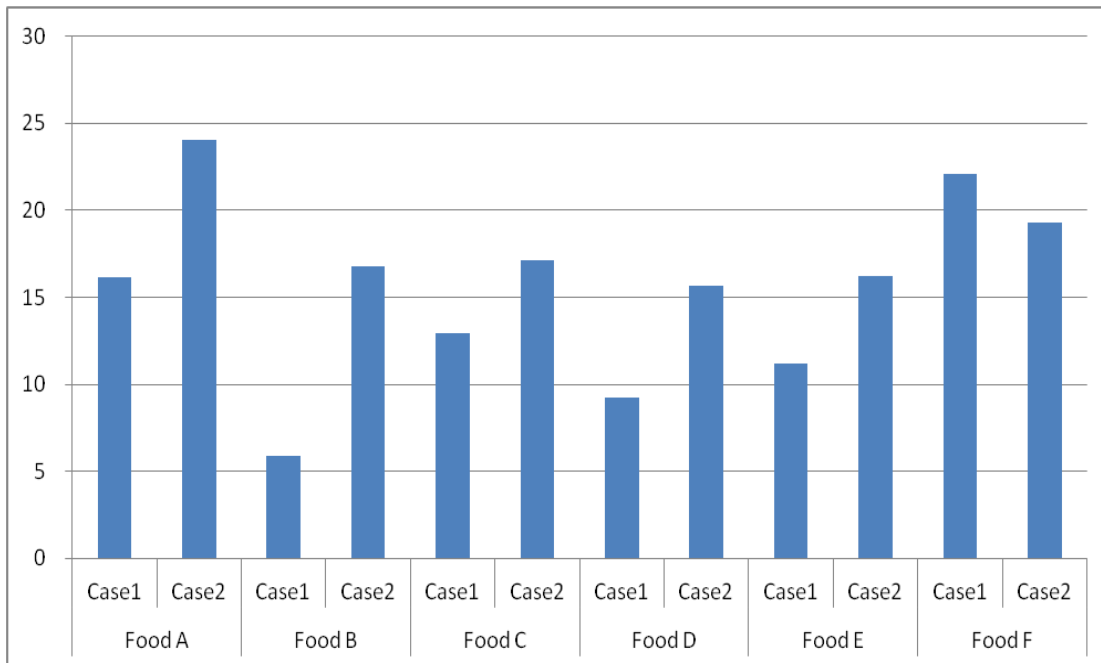


Fig. 5-19. FI



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VI. CONCLUSIONS

Focusing on the idea that the equation of exponential smoothing method (ESM) was equivalent to (1,1) order ARMA model equation, a new method of estimation of smoothing constant in exponential smoothing method was proposed before by us which satisfied minimum variance of forecasting error. Generally, smoothing constant was selected arbitrarily. But in this paper, we utilized above stated theoretical solution. Firstly, we made estimation of ARMA model parameter and then estimated smoothing constants. Thus theoretical solution was derived in a simple way and it might be utilized in various fields.

Furthermore, combining the trend removal method with this method, we aimed to increase forecasting accuracy. An approach to this method was executed in the following method. Trend removal by a linear function was applied to the original sales data of a shop. The combination of linear and non-linear function was also introduced in trend removing. "a day of the week index (DWI)" is newly introduced for the daily sales data and a day of the week trend is removed. Theoretical solution of smoothing constant of ESM was calculated for both of the DWI trend removing data and the non DWI trend removing data. Then forecasting was executed on these data.

In 5 cases out of 6, case 1 (DWI is imbedded) is better than case 2 (DWI is not imbedded). In particular, Food B obtained more than double twice improvement in forecasting accuracy. Food D had also the good result. The effectiveness of this method should be examined in various cases.

REFERENCES

- [1] Box Jenkins: Time Series Analysis Third Edition (Prentice Hall, 1994)
- [2] F.R.Johnston: Exponentially Weighted Moving Average (EWMA) with Irregular Updating Periods. Journal of the Operational Research Society Vol. 44, No. 7 (1993), 711-716.
- [3] Hidekatsu Tokumaru et al.: Analysis and Measurement –Theory and Application of Random data Handling (Baifukan Publishing, 1982)
- [4] Katsuro Maeda: Smoothing Constant of Exponential Smoothing Method. Seikei University Report Faculty of Engineering No. 38 (1984), 2477-2484.
- [5] Kazuhiro Takeyasu: System of Production, Sales and Distribution (Chuoikeizai-Sha Publishing, 1996)
- [6] Kazuhiro Takeyasu and Kazuko Nagao: Estimation of Smoothing Constant of Minimum Variance and its Application to Industrial Data. Industrial Engineering and Management Systems Vol. 7, No. 1 (2008), 44-50.
- [7] Kengo Kobayashi: Sales Forecasting for Budgeting (Chuoikeizai-Sha Publishing, 1992)
- [8] M.West and P.J.Harrison: Bayesian Forecasting and Dynamic Models (Springer-Verlag, New York, 1989)
- [9] Naohiro Ishii et al.: Bilateral Exponential Smoothing of Time Series. Int.J.System Sci. Vol. 12, No. 8 (1991), 997-988.
- [10] Peter R.Winters: Forecasting Sales by Exponentially Weighted Moving Averages. Management Science Vol. 6, No. 3 (1984), 324-343.
- [11] R.G. Brown: Smoothing, Forecasting and Prediction of Discrete – Time Series (Prentice Hall, 1963)
- [12] Spyros Makridakis and Robeat L.Winkler: Averages of Forecasts Some Empirical Results. Management Science Vol. 29, No. 9 (1983), 987-996.
- [13] Steinar Ekern: Adaptive Exponential Smoothing Revisited. Journal of the Operational Research Society Vol. 32 (1982), 775-782.

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